

Evaluating the Dimensionality of the Michigan English Language Assessment Battery

Hong Jiao
Florida State University

This study evaluates the dimensionality of Form FF, listening and Form EE, grammar/cloze/vocabulary/reading (GCVR) in the Michigan English Language Assessment Battery (MELAB). It further investigates the influences of gender, native language, and proficiency level on the dimensionality of the listening and the GCVR sections in the MELAB. Stout's procedure was employed to test two hypotheses; that the listening items are unidimensional, and the GCVR items are unidimensional. Principle axes factor analysis and principle component analysis both using the tetrachoric correlation matrix were applied for further exploration. The results indicate that both listening and GCVR tests were unidimensional for female and Tagalog/Filipino-speaking groups. Further, the global GCVR test was unidimensional. But for other groups, the results were inconsistent across methods regarding the unidimensionality of both forms.

Validity is an important issue in test development and evaluation according to the *Standards for Educational and Psychological Testing* (American Educational Research Association, American Psychological Association, & National Council on Measurement in Education, 1999), as well as fairness. Validity refers to the degree to which evidence supports the inferences based on test scores (Messick, 1989). Fairness means that all examinees are given comparable opportunities to demonstrate their abilities on the construct a test intends to measure (American Educational Research Association, et al., 1999, p.74). In examining test fairness, the researcher should address questions such as whether the test measures the same construct in all relevant populations (Wang & Witt, 2002). An investigation of the factor structure of a test can provide evidence of validity (Messick, 1995) and fairness of the test.

To make valid and fair comparisons across examinee groups, test items should be constructed to measure the same construct(s). However, test items often measure other traits in addition to the traits they intend to measure (Hambleton & Swaminathan, 1985; Reckase, 1979, 1985; Stout, 1987). Several studies (Birenbaum & Tatsuoka, 1982; Bock, Gibbons, & Muraki, 1988) indicated that item attributes and examinee characteristics could affect the dimensionality of test items. Dimensionality is a property of the test as well as of the examinees (Reckase, 1990). This is also true with language proficiency tests. Oltman and Stricker (1988) and Kok (1992) showed that examinees' English proficiency and ethnic background affected the dimensionality of language proficiency tests. Therefore, when evaluating the dimensionality of a test, it would provide a better picture if the evaluation were done from the perspective of both the test itself and the examinees' characteristics.

The Michigan English Language Assessment Battery (MELAB) is a test developed for college and university admission. It intends to assess the English language competence of adult non-native speakers of English who will apply to study at an English-speaking institution. The MELAB includes Part 1, writing on a prompt, Part 2, a listening

comprehension test, Part 3, grammar, cloze, vocabulary, and reading comprehension (GCVR) questions, and an optional speaking test. All the items in the listening and GCVR subtests are scored dichotomously.

The *MELAB Technical Manual* (English Language Institute, University of Michigan, 1996) reported factor analysis results, done at the component score level, of Part 2 and Part 3 in the MELAB. The *Michigan English Language Assessment Battery: Technical Manual 2003* (English Language Institute, University of Michigan, 2003) reported an item factor analysis of the individual items in Part 2 and Part 3 sections of the MELAB, and principle component analysis of the testlet scores. The item-level factor analysis provided more information regarding the construct validity of Part 2 and Part 3 in the MELAB.

The purpose of this study is to evaluate the dimensionality of Part 2, listening comprehension, and Part 3, GCVR, in the MELAB separately, using item-level information. It further investigates the influences of gender, native language, and language proficiency level on the dimensionality of the listening comprehension test and the GCVR test. Two sets of hypotheses were proposed for this research. One set was that the items in Part 2 were unidimensional, globally and across subgroups of examinees. The other set was that the items in Part 3 were also unidimensional, globally and across subgroups of examinees. In short, this study intends to provide a better understanding of the dimensionality of Part 2 and Part 3 in the MELAB and to determine the consistency in dimensionality across all available examinees and the selected subgroups of examinees.

Stout's Nonparametric Analysis of Dimensionality

Stout (1987) proposed the concept of essential unidimensionality. Essential unidimensionality refers to the existence of exactly one dominant dimension. Van Abswoude, van der Ark, and Sijtsma (2004) suggest that DIMTEST can be used to verify unidimensionality when there is only one dominant trait in the data set and only a few items are driven by another trait. DIMTEST is a statistical assessment of whether there is one or more than one dominant dimension. Stout's nonparametric procedure (Nandakumar & Stout, 1993; Stout, 1987) tests two statistical hypotheses. The null hypothesis states that $d = 1$ and the alternative hypothesis is that $d > 1$, where d stands for the number of dimensions in a set of test items.

To test the hypothesis that a group of dichotomous items is unidimensional, DIMTEST divides all the items on the test into three subtests, two assessment subtests, and one partitioning subtest. DIMTEST selects M items from N items: the total number of items on the test. These M items are called Assessment Subtest 1, or AT1. These items are dimensionally distinct from the rest of the items and measure the same trait. AT1 items can be selected based on expert opinions or based on factor analysis results. DIMTEST can automatically choose the AT1 items through a factor analysis procedure using the tetrachoric correlation matrix. Then, another group of M items are selected from the remaining items so that the difficulty levels of these items are as similar to those of the AT1 items as possible. They are called Assessment Subtest 2, AT2. AT2 has a similar difficulty distribution as does AT1, and is dimensionally similar to the remaining items. AT2 is constructed to reduce the examinee variability bias and the item difficulty bias, which may result in false rejection of the null hypothesis. The remaining items, $n = N - 2M$, is the Partitioning Subtest, PT. Examinees are divided into subgroups based on their PT scores. Examinees with the same PT

score are assigned into the same PT subgroup. For each PT subgroup, a statistic, t_{Lk} , based on the examinees' AT1 subtest scores, and a statistic, t_{Bk} , based on the examinees' AT2 subtest scores are calculated (see Nandakumar & Stout, 1993; Stout, 1987, for details). The DIMTEST statistic T is obtained with the following formula:

$$T = \frac{T_L - T_B}{\sqrt{2}}$$

where

$$T_L = \sum_{k=1}^{n-1} t_{Lk} \quad \text{and} \quad T_B = \sum_{k=1}^{n-1} t_{Bk}$$

(k is the number of PT scores, t_{Lk} is the standardized difference between two variance estimates). The first variance estimate is actual observed variance between subgroups k number-correct scores on AT1. The second is the variance between subgroups k number-correct AT1 scores that were predicted under the unidimensional assumption. The first variance estimate will be inflated if the data is multidimensional. The measure of the amount of multidimensionality for subgroup k is t_{Lk} . If unidimensionality holds for subgroup k except for statistical error, the two within-subgroup variance estimates are approximately equal, and $t_{Lk} = 0$. If multidimensionality holds, then $t_{Lk} > 0$. If the test is not long enough, T_L is biased even if unidimensionality holds. Thus, T_B is used to correct the statistical bias in T_L . It is better to apply DIMTEST for tests with the total number of items equal to or larger than 80 items. Van Abswoude, et al. (2004) indicated that DIMTEST has low power for short tests.

Consequently, Stout's statistic, T , is compared with the upper $100(1-\alpha)$ percentile of the standard normal distribution, Z_α , for a desired level of significance, α . When $T < Z_\alpha$, DIMTEST accepts the unidimensionality hypothesis. When $T > Z_\alpha$, DIMTEST rejects the unidimensionality hypothesis. Nandakumar (1991) modified Stout's more conservative statistic and proposed a more powerful statistic with a slightly higher but acceptable type I error rate. This study presents both Stout's conservative statistic T and Nandakumar's more powerful statistic T' .

Since DIMTEST uses test scores as a conditioning variable, it shows some positive bias even after correcting for the two types of bias using AT2. To further reduce bias and increase the power of the more powerful statistic T' , Stout, Goodwin Froelich, and Gao (2001) proposed a new DIMTEST procedure that uses only one AT subtest.

DIMTEST incorporates a FAC procedure that performs an unrotated principle axis factor analysis of the tetrachoric correlation matrix for the dichotomous data set with maximum interitem correlations estimating communalities. A FAC output file containing the second-factor loadings can be utilized as an input file for the ASN program in automatically selecting AT1 items. When the ASN program uses the second-factor loadings from the output of FAC program to automatically choose AT1 items, the number of AT1 items is also determined automatically to optimize the statistical power (Nandakumar & Stout, 1993). The other output file from the FAC program contains the results of the factor analysis. AT1 items can be selected subjectively based on the FAC output. Three sub-programs in DIMTEST are used in this study: FAC, ASN, and SSC. FAC implements a tetrachoric factor analysis, ASN uses the output from the FAC program to select AT1 and AT2 items, and SSC calculates the DIMTEST statistics.

Principle Component Analysis with Tetrachoric Correlations

Principle component analysis has been used by researchers to assess dimensionality of a set of items (Abedi, 1997). If a large amount of variance can be explained by the first component, the set of items can be considered unidimensional. For dichotomous items, principle components analysis uses tetrachoric correlations to reduce items into a small number of principle components accounting for most of the variance in the items. According to Hatcher (1994), multiple criteria can be used to determine the number of components to retain. They are the eigenvalue, the proportion of variance accounted for, and the interpretability criterion. Kaiser (1960) suggests that a component with an eigenvalue larger than one be retained. Hatcher (1994) suggests retaining any component accounting for at least 10% of the total variance. Reckase (1979) suggests that if the first component explains 20% of the variance of a set of items, the item set is unidimensional. In addition, the interpretation of the retained component should make substantive meaning of the constructs.

Method

Data

The data used in this study were from examinees who took Form EE of the Part 2 (GCVR) and Form FF of the Part 3 (listening) of the MELAB. Their responses to Part 2 and Part 3, and the information on their gender and native language, were collected. There were 1,031 examinees who took Form EE, and 1,650 examinees who took Form FF. Form EE contains 100 items: Items 1 through 30 are grammar items, items 31 through 50 are cloze passage items, items 51 through 80 measure vocabulary, and items 81 through 100 are the reading comprehension items. For Form FF (listening), items 1 through 15 are the short questions, items 16 through 35 are the short conversations, and items 36 through 50 are radio report comprehension items. Table 1 shows the test composition for Form EE and Form FF.

Table 1. Composition for Form EE and Form FF

Form EE (GCVR)	Item Numbers	Form FF (Listening)	Item Numbers
Grammar	1 to 30	Short questions	1 to 15
Cloze	31 to 50	Short conversations	16 to 35
Vocabulary	51 to 80	Radio reports	36 to 50
Reading	81 to 100		

In this study, two sets of analyses were carried out. The first set of analyses focuses on the global structure of the test items using all available examinee data, and the second set evaluates the local structure of data across subgroups of test takers that differ with respect to gender, native language, and proficiency level. The selected six sub-groups of examinees were female, male, examinees whose native language was Korean, those whose native language was Tagalog/Filipino, high-proficiency examinees for a particular subtest, and low-proficiency examinees for a particular subtest. For both the listening and GCVR tests, high and low proficiency levels were distinguished by the examinee scale scores. For both subtests, examinees with a scale score of 80 or higher were classified as high proficiency

examinees, and those with a scale score below 80 were categorized as low proficiency examinees. Due to the limited amount of data, only Tagalog/Filipino and Korean language groups were analyzed. Altogether, seven dichotomous response data sets were set up for each subtest; namely, data sets for all available examinees taking a particular subtest, female examinees, male examinees, Tagalog/Filipino speakers, Korean speakers, high-proficiency examinees, and low-proficiency examinees, resulting in 14 data sets. The number of examinees for the seven groups for each subtest is summarized in Table 2.

Table 2. Number of Examinees for Each Data Set

Listening	Number of Examinees
All examinees	1,650
Female	1,257
Male	393
Tagalog/Filipino	800
Korean	141
High listening proficiency	939
Low listening proficiency	711
GCVR	
All examinees	1,031
Female	787
Male	244
Tagalog/Filipino	527
Korean	123
High GCVR proficiency	528
Low GCVR proficiency	503

Procedure

After the collection of item response data, multiple procedures were adopted to appraise global and local dimensionality of the tests: Stout's DIMTEST procedure, principle axis factoring of tetrachoric correlations, and principle components analysis with LISREL.

To calculate Stout's statistic, test items were split into three subtests: AT1 (assessment subtest 1), AT2 (assessment subtest 2), and PT (partitioning subtest). The items in the AT1 were not specified in advance, but were identified through a principle axis factor analysis of the tetrachoric correlations by the FAC application. The FAC program first ran the tetrachoric factor analysis to calculate the second-factor loadings for the selection of AT1 subtest items set. Only part of the data was used for running FAC, while the rest of the data was used for ASN and SSC runs. The ASN program selected AT1 and AT2 items automatically based on the FAC output, and the SSC program calculated the DIMTEST statistic. SSC assessed the statistical significance of the distinctiveness of the dimensionality between two specified subtests: the Assessment Subtests and the Partitioning Subtest. DIMTSET conservative and more powerful T statistics, and their probabilities, were used as the criteria to evaluate test dimensionality, and $\alpha = 0.05$ was used to determine the significance of the hypothesis test.

Principle axis factoring of tetrachoric correlations was run for each group of examinees. Tetrachoric factor analysis in the DIMTEST was run with the specification of the number of factors. For this study, the number of factors was determined based on the natural groupings of the items. For Form FF, the 50 listening items were divided into three parts, measuring examinee ability to understand short questions, short conversations, and radio reports. For Form EE, the 100 GCVR items were categorized into 4 groups, measuring examinee proficiency in grammar, cloze, vocabulary, and reading comprehension. The resulting number of factors for the tetrachoric factor analysis was determined based on the criteria of the strength of the eigenvalues and the differences between the factor eigenvalues.

To further confirm the results obtained in the principle axis factor analysis and the DIMTEST procedure, LISREL was employed to run a principle component analysis using the tetrachoric correlation matrix. The percentage of variance explained by each factor was employed in determining the number of principle components.

Stout's nonparametric procedure, tetrachoric factor analysis, and principle component analysis using the tetrachoric correlation matrix, were each applied to each of the seven data sets for both test forms. Form EE and Form FF of the MELAB test were analyzed for all available examinees and the six specified subgroups of examinees using DIMTEST, FAC in the DIMTEST, and LISREL. Results based on all the available data and those based on the data from the subgroups were compared to evaluate the global dimensionality and the consistency of local dimensionality across different subgroups of examinees.

Results

Stout's DIMTEST Statistics

Table 3 summarizes the analysis results from Stout's nonparametric procedure to test the unidimensionality of the 14 data sets using DIMTEST. For the listening test, both conservative and the more powerful tests indicated that the test was not unidimensional based on all available examinee responses and for the male subgroup of examinees. However, the listening test was unidimensional at the 0.05 level for the female, Tagalog/Filipino speakers, and both high and low listening proficiency examinee groups. For examinees whose native language was Korean, the listening test was unidimensional based on the conservative statistics, but not unidimensional based on the more powerful statistic.

For the GCVR test, the unidimensionality test was accepted at the 0.05 level for all examinees, females, Tagalog/Filipino speakers, and both high- and low-proficiency subgroups. For the male examinees, the GCVR test was not unidimensional. For the Korean-speaking examinees, the DIMTEST conservative statistic supported unidimensionality, but the more powerful statistic rejected the null hypothesis that the GCVR test was unidimensional.

Table 3. DIMTEST Results

	DIMTEST Statistics			
	Conservative (<i>T</i>)		More Powerful (<i>T</i>)	
	<i>T</i>	P-value	<i>T</i>	P-value
<i>Listening</i>				
All examinees	1.991	0.023	2.254	0.012
Female	0.805	0.210	1.034	0.151
Male	2.625	0.004	3.108	0.001
Tagalog/Filipino	0.529	0.298	0.697	0.243
Korean	1.457	0.073	1.943	0.026
High listening proficiency	0.282	0.389	0.404	0.343
Low listening proficiency	0.325	0.373	0.425	0.335
<i>GCVR</i>				
All examinees	-0.500	0.692	-0.473	0.682
Female	1.088	0.138	1.421	0.078
Male	1.802	0.036	2.148	0.016
Tagalog/Filipino	-3.441	0.9997	-4.159	0.99998
Korean	1.288	0.099	2.176	0.015
High GCVR proficiency	-0.413	0.66	-0.485	0.686
Low GCVR proficiency	0.071	0.472	-0.004	0.502

Eigenvalues and Eigenvalue Differences

The eigenvalues and the eigenvalue differences between factors for the listening test are summarized in Table 4. The eigenvalues for the first three factors were all larger than 1 for all groups. For all examinees, female, male, Tagalog/Filipino, and Korean-speaking examinee groups, the eigenvalues for the first factor were larger than 10, and the difference between the first two factors was around 10. Hattie (1985) suggests using the difference between the first factor and the second factor divided by the difference between the second and the third factor to examine the relative strength of the first factor (dubbed the Factor Difference Ratio Index (FDRI) in Johnson, Yamashiro, & Yu, 2003). If this ratio is larger than 3, the first factor is relatively strong. Based on this criterion, the data for six examinee groups: all examinees, female, male, Tagalog/Filipino speakers, Korean speakers, and low listening proficiency, satisfy this criterion. The first factor was relatively strong for these six groups of examinees. For the examinee group with high listening proficiency, the eigenvalues for the three factors were of similar strength, around 5. This data set did not meet the FDRI criterion suggested by Hattie (1985).

The eigenvalues and the eigenvalue differences between factors for the GCVR data are summarized in Table 5. The eigenvalues for the first four factors were larger than 1 for all groups. The eigenvalues for the first factor was larger than 10 for all groups, but the FDRI was greater than 3 for only all examinees, female, male, Tagalog/Filipino, and low-GCVR-proficiency examinee groups. The examinee group with high GCVR proficiency and the Korean subgroup did not satisfy Hattie's (1985) FDRI recommendation.

Table 4. Eigenvalues from Tetrachoric Factor Analysis — Listening

Listening	Factor	Eigenvalue	Difference	FDRI
All examinees	1	12.133	10.477	22.68
	2	1.655	0.462	
	3	1.193	0.182	
Female	1	11.994	10.297	23.83
	2	1.698	0.431	
	3	1.266	0.153	
Male	1	12.810	9.917	11.17
	2	2.893	0.888	
	3	2.005	0.306	
Tagalog/Filipino	1	11.914	10.086	20.80
	2	1.828	0.485	
	3	1.343	0.218	
Korean	1	13.903	10.217	17.32
	2	3.685	0.590	
	3	3.095	0.414	
High listening proficiency	1	5.993	1.153	1.63
	2	4.839	0.708	
	3	4.132	0.674	
Low listening proficiency	1	5.602	3.996	26.29
	2	1.606	0.152	
	3	1.454	0.246	

FDRI = factor difference ratio index $((F1-F2)/(F2-F3))$.

Table 5. Eigenvalues from Tetrachoric Factor Analysis — GCVR

GCVR	Factor	Eigenvalue	Difference	FDRI
All examinees	1	26.423	20.510	5.79
	2	5.912	3.542	
	3	2.370	0.291	
	4	2.079	0.576	
Female	1	25.147	18.838	5.21
	2	6.308	3.616	
	3	2.692	0.467	
	4	2.225	0.430	
Male	1	30.771	25.005	11.03
	2	5.767	2.267	
	3	3.500	0.807	
	4	2.693	0.171	
Tagalog/Filipino	1	22.959	16.987	16.92
	2	5.972	1.004	
	3	4.968	0.342	
	4	4.626	0.681	

Table 5. (cont.)

GCVR	Factor	Eigenvalue	Difference	FDRI
Korean	1	16.148	6.426	1.40
	2	9.722	4.603	
	3	5.119	0.776	
	4	4.343	0.351	
High GCVR proficiency	1	11.886	3.427	2.69
	2	8.460	1.272	
	3	7.187	2.118	
	4	5.069	0.696	
Low GCVR proficiency	1	11.729	7.176	4.17
	2	4.553	1.719	
	3	2.834	0.180	
	4	2.654	0.726	

FDRI = factor difference ratio index $((F1-F2)/(F2-F3))$.

Percentage of Variance Explained

To further check the results from the tetrachoric factor analysis in DIMTEST, principle component analysis using the tetrachoric correlation matrix was run in LISREL. The analysis results from LISREL are summarized in Table 6 for the listening test and Table 7 for the GCVR test. For the eigenvalue, a similar pattern to that based on DIMTEST tetrachoric factor analysis results was observed for both listening and GCVR tests using LISREL. Using Reckase's (1979) suggestion that a set of test items is unidimensional if the first factor accounts for 20% or more of the total variance, for the listening test, the item responses of all examinees, females, males, Tagalog/Filipino speakers, and Korean speakers are unidimensional. However, the first principle component for the high and low listening proficiency groups accounted for much less than 20% of the total variance.

For the GCVR test, the first factor for the examinee groups of all examinees, female, male, and Tagalog/Filipino speakers accounted for more than 20% of the total variance. For the Korean speakers and high- and low-GCVR-proficiency examinee groups, the first principle component accounted for less than 20 % of the total variance.

The results regarding the dimensionality of the data sets for the listening and the GCVR tests in the MELAB are summarized in Table 8. In most cases, different procedures and evaluation criteria led to different conclusions regarding the dimensionality of the test data. Only data sets from 5 out of 14 groups of examinees yielded consistent conclusions. The listening test for female and Tagalog/Filipino-speaking groups, and the GCVR test for female, Tagalog/Filipino speakers, and all examinees were unidimensional consistently across all procedures used in this study.

Summary and Discussion

Regarding the global dimensionality of Form EE (GCVR test) in the MELAB, unidimensionality held. This is consistent with the item-level factor analysis results reported in the MELAB technical manual 2003 for Part 3. On the other hand, inconsistent results were obtained regarding the global unidimensionality of Form FF (listening test).

Table 6. Eigenvalues from LISREL – Listening

Listening		PC 1	PC 2	PC 3
All examinees	Eigenvalue	12.69	2.26	1.80
	% Variance	25.38	4.52	3.60
	Cum. % Var	25.38	29.89	33.49
Female	Eigenvalue	12.55	2.30	1.87
	% Variance	25.10	4.59	3.73
	Cum. % Var	25.10	29.69	33.42
Male	Eigenvalue	13.24	3.03	2.24
	% Variance	26.47	6.07	4.48
	Cum. % Var	26.47	32.54	37.02
Tagalog/Filipino	Eigenvalue	12.45	2.39	1.94
	% Variance	24.90	4.79	3.88
	Cum. % Var	24.90	29.69	33.57
Korean	Eigenvalue	14.15	3.79	3.15
	% Variance	28.30	7.58	6.31
	Cum. % Var	28.30	35.89	42.19
High listening proficiency	Eigenvalue	4.81	4.14	3.46
	% Variance	9.62	8.28	6.39
	Cum. % Var	9.62	17.91	24.83
Low listening proficiency	Eigenvalue	6.18	2.33	2.18
	% Variance	12.36	4.65	4.37
	Cum. % Var	12.36	17.01	21.38

Table 7. Eigenvalues from LISREL — GCVR

GCVR		PC 1	PC 2	PC 3	PC 4
All examinees	Eigenvalue	26.90	6.41	2.84	2.56
	% Variance	26.90	6.41	2.84	2.56
	Cum. % Var	26.90	33.32	36.15	38.72
Female	Eigenvalue	25.61	6.81	3.17	2.78
	% Variance	25.61	6.81	3.17	2.78
	Cum. % Var	25.61	32.42	35.59	38.37
Male	Eigenvalue	31.08	6.13	3.60	3.09
	% Variance	31.08	6.13	3.60	3.09
	Cum. % Var	31.08	37.21	40.81	43.90
Tagalog/Filipino	Eigenvalue	22.03	7.57	6.39	4.35
	% Variance	22.03	7.57	6.39	4.35
	Cum. % Var	22.03	29.60	35.99	40.34

Table 7. (cont.)

GCVR		PC 1	PC 2	PC 3	PC 4
Korean	Eigenvalue	16.49	10.07	5.48	4.76
	% Variance	16.49	10.07	5.48	4.76
	Cum. % Var	16.49	26.56	32.05	36.80
High GCVR proficiency	Eigenvalue	12.35	8.86	6.32	4.73
	% Variance	12.35	8.86	6.32	4.73
	Cum. % Var	12.35	21.21	27.53	32.27
Low GCVR proficiency	Eigenvalue	12.26	5.18	3.50	3.30
	% Variance	12.26	5.18	3.50	3.30
	Cum. % Var	12.26	17.43	20.93	24.23

Table 8. Results Summary

	DIMTEST statistics		Eigenvalue difference	Percentage of variance explained by first factor
	Conservative	More powerful		
Listening				
All examinees	N	N	U	U
Female	U	U	U	U
Male	N	N	U	U
Tagalog/Filipino	U	U	U	U
Korean	U	N	U	U
High listening proficiency	U	U	N	N
Low listening proficiency	U	U	U	N
GCVR				
All examinees	U	U	U	U
Female	U	U	U	U
Male	N	N	U	U
Tagalog/Filipino	U	U	U	U
Korean	U	N	N	N
High GCVR proficiency	U	U	N	N
Low GCVR proficiency	U	U	U	N

For each of the studied subgroups of examinees, unidimensionality held for both forms for female and Tagalog/Filipino-speaking examinee groups. For other subgroups of examinees, the results regarding the unidimensionality for each of the two forms were inconsistent in terms of the evaluation criteria adopted in this study. Thus, no conclusion can be drawn regarding the dimensionality of the listening and GCVR sections of the test for those groups of examinees. The dimensionality was not consistent across different groups of examinees for each test form.

Several issues need attention regarding the generalizability of the study results. The first issue is related to the sample size in this study. Hatcher (1994) suggested that the minimal sample size should be 5 times the number of variables to be analyzed. Van Abswoude, et al. (2004) suggested that DIMTEST may be more effective for large sample sizes, such as $n = 2,000$. However, there were only data for 1,031 examinees available for the GCVR test, and 1,650 examinees for the listening test. When the subgroups of examinees were selected from these samples, much smaller samples resulted. For example, for the male and Korean groups, there were fewer than 400 cases. Sample sizes smaller than that desired may affect the analysis results.

Stout's nonparametric procedure requires selection of the AT1 items, which was completed automatically by the program based on the second-factor loadings and other output from the FAC analysis. This study also tried to manually select AT1 items based on the tetrachoric factor analysis results. However, such selection may lead to the failure of the Wilcoxon Rank Sum Test in the ASN program. Thus, expert opinions and the factor loadings from the tetrachoric factor analysis presented in Table A.1 to Table B.7 in Appendices A and B may be combined to come up with a set of AT1 items. This may improve the DIMTEST procedure for assessing dimensionality.

In summary, the results of this study provide more information regarding the dimensionality of Part 2, listening, and Part 3, GCVR, in the MELAB. This study helps to identify the effect of the test items and of the examinees' characteristics on the dimensionality of the MELAB sections.

Acknowledgments

I would like to express my gratitude to Dr. Shudong Wang for his insightful comments and valuable suggestions. My thanks also go to Xueli Xu for her help on the DIMTEST program. Finally, I would like to sincerely thank Jeff Johnson for his assistance in securing the MELAB data, the information on the tests, proofreading and editing the paper.

References

- Abedi, J. (1997). *Dimensionality of NAEP subscale scores in mathematics* (CSE Tech. Rep. 428). Los Angeles, CA: CRESST/University of California, Los Angeles.
- American Educational Research Association, American Psychological Association, & National Council on Measurement in Education. (1999). *Standards for educational and psychological testing*. Washington, DC: American Educational Research Association.
- Birenbaum, M., & Tatsuoka, K. K. (1982). On the dimensionality of achievement data. *Journal of Educational Measurement, 19*(4), 259-266.
- Bock, R. D., Gibbons, R., & Muraki, E. (1988). Full-information item factor analysis. *Applied Psychological Measurement, 12*(3), 261-280.
- English Language Institute, University of Michigan. (1996). *MELAB technical manual*. Ann Arbor, MI: English Language Institute, University of Michigan.
- English Language Institute, University of Michigan. (2003). *Michigan English Language Assessment Battery: Technical manual 2003*. Ann Arbor, MI: English Language Institute, University of Michigan.

- Hambleton, R. K., & Swaminathan, H. (1985). *Item response theory: Principles and applications*. Boston: Kluwer Nijhoff.
- Hatcher, L. (1994). *A step-by-step approach to using the SAS system for factor analysis and structural equation modeling*. Cary, NC: SAS Institute Inc.
- Hattie, J. (1985). Methodology review: Assessing unidimensionality of tests and items. *Applied Psychological Measurement*, 9(2), 139-164.
- Johnson, J. S., Yamashiro, A., & Yu, J. (2003). *ECPE annual report: 2002*. Ann Arbor, MI: English Language Institute, University of Michigan.
- Kaiser, H. F. (1960). The application of electronic computers to factor analysis. *Educational and Psychological Measurement*, 20, 141-151.
- Kok, F. (1992). Differential item functioning. In L. Verhoeven, and J. H. A. L. de Jong, (Eds.), *The construct of language proficiency: Applications of psychological models to language assessment* (pp.115-124). Amsterdam: John Benjamins.
- Messick, S. (1989). Validity. In R.L. Linn (Ed.), *Educational measurement* (3rd ed., pp. 13 - 103). New York: American Council on Education and Macmillan.
- Messick, S. (1995). Standards-based score interpretation: Establishing valid grounds for valid inferences. In *Proceedings of the joint conference on standard setting for large-scale assessments of the National Assessment Governing Board (NAGB) and the National Center for Education Statistics (NCES), Vol. II* (pp. 291-305). Washington, DC: National Assessment Governing Board and National Center for Education Statistics.
- Nandakumar, R. (1991). Traditional dimensionality versus essential dimensionality. *Journal of Educational Measurement*, 28(2), 99-118.
- Nandakumar, R., & Stout, W. (1993). Refinements of Stout's procedure for assessing latent trait unidimensionality. *Journal of Educational Statistics*, 18(1), 41-68.
- Oltman, P. K., & Stricker, L., J. (1988, March). *How native language and level of English proficiency affect the structure of the Test of English as a Foreign Language*. Paper presented at the Annual Colloquium on Language Testing Research, Urbana, IL.
- Reckase, M. D. (1979). Unifactor latent trait models applied to multifactor tests: Results and implications. *Journal of Educational Statistics*, 4(3), 207-230.
- Reckase, M. D. (1985). The difficulty of test items that measure more than one ability. *Applied Psychological Measurement*, 9(4), 401-412.
- Reckase, M. D. (1990, April). *Unidimensionality data from multidimensional tests and multidimensional data from unidimensional tests*. Paper presented at the Annual Meeting of the American Educational Research Association, Boston, MA.
- Stout, W. F. (1987). A nonparametric approach for assessing latent trait unidimensionality. *Psychometrika*, 52(4), 589-617.
- Stout, W. F., Goodwin Froelich, A., & Gao, F. (2001). Using resampling methods to produce an improved DIMTEST procedure. In A. Boomsma, M. A. J. van Duijn, & T. A. B. Snijders (Eds.), *Essays on item response theory* (pp. 357-375). New York: Springer.
- van Abswoude, A. A. H., van der Ark, L. A., & Sijtsma, K. (2004). A comparative study of test data dimensionality assessment procedures under nonparametric IRT models. *Applied Psychological Measurement*, 28(1), 3-24.
- Wang, S., & Witt, E. (2002, April). *Validation and invariance of factor structure of a national licensing examination across gender and race*. Paper presented at the annual meeting of the American Educational Research Association, New Orleans, LA.

Appendix A

Factor Loadings for Form FF-Listening

Table A.1. Factor Loadings for Listening-All Examinees

Item	Factor Loadings		
	1	2	3
1	0.611	-0.078	-0.162
2	0.376	0.056	-0.199
3	0.687	-0.137	-0.122
4	0.433	0.076	0.086
5	0.435	0.024	-0.221
6	0.582	-0.161	-0.218
7	0.645	-0.088	-0.186
8	0.544	0.041	-0.101
9	0.547	0.097	-0.284
10	0.330	-0.051	-0.056
11	0.244	-0.015	-0.149
12	0.566	-0.153	0.063
13	0.386	0.104	-0.146
14	0.407	0.133	-0.080
15	0.555	0.124	-0.111
16	0.579	-0.348	0.041
17	0.686	-0.054	-0.044
18	0.484	-0.096	0.093
19	0.572	-0.153	0.187
20	0.655	-0.002	0.014
21	0.413	-0.052	-0.108
22	0.665	-0.239	0.090
23	0.737	-0.147	-0.058
24	0.530	-0.084	0.141
25	0.326	-0.153	0.219
26	0.576	-0.110	-0.024
27	0.424	0.066	-0.243
28	0.701	-0.192	0.179
29	0.696	-0.226	-0.015
30	0.368	-0.005	-0.226
31	0.481	0.070	-0.022
32	0.513	-0.216	0.134
33	0.485	0.184	-0.091
34	0.512	0.305	-0.113
35	0.312	0.048	0.063
36	0.515	-0.037	0.273
37	0.383	0.279	0.081
38	0.314	0.159	0.011
39	0.098	0.110	0.140
40	0.535	0.054	-0.095

41	0.360	0.240	0.176
42	0.339	0.280	0.145
43	0.449	0.007	0.372
44	0.515	0.163	0.215
45	0.495	0.158	0.322
46	0.259	0.502	0.012
47	0.418	0.243	0.128
48	0.270	0.166	0.118
49	0.260	0.518	-0.072
50	0.332	0.198	0.068

Table A.2. Factor Loadings for Listening-Female

Item	Factor Loadings		
	1	2	3
1	0.620	-0.079	-0.158
2	0.393	0.006	-0.202
3	0.658	-0.176	-0.076
4	0.441	0.056	0.059
5	0.431	-0.009	-0.139
6	0.574	-0.200	-0.221
7	0.668	-0.053	-0.165
8	0.528	0.011	-0.133
9	0.517	0.049	-0.308
10	0.350	0.020	-0.036
11	0.210	-0.121	-0.168
12	0.564	-0.126	-0.011
13	0.390	0.155	-0.175
14	0.399	0.075	-0.100
15	0.547	0.074	-0.143
16	0.580	-0.365	0.122
17	0.678	-0.124	-0.019
18	0.497	-0.087	0.104
19	0.565	-0.132	0.153
20	0.611	-0.009	0.103
21	0.427	-0.004	-0.197
22	0.642	-0.190	0.043
23	0.726	-0.128	-0.064
24	0.531	-0.007	0.128
25	0.352	-0.109	0.202
26	0.605	-0.143	0.035
27	0.388	0.067	-0.248

28	0.704	-0.196	0.200
29	0.706	-0.254	0.005
30	0.375	-0.038	-0.200
31	0.477	0.004	0.017
32	0.510	-0.119	0.095
33	0.499	0.204	-0.147
34	0.507	0.344	-0.187
35	0.320	0.015	0.068
36	0.502	-0.046	0.344
37	0.373	0.286	0.091
38	0.324	0.179	0.052
39	0.112	0.098	0.080
40	0.515	0.107	-0.111
41	0.367	0.214	0.206
42	0.340	0.300	0.189
43	0.460	0.056	0.341
44	0.531	0.162	0.179
45	0.474	0.187	0.323
46	0.221	0.515	-0.023
47	0.410	0.318	0.104
48	0.279	0.191	0.087
49	0.245	0.478	-0.127
50	0.345	0.187	0.052

Table A.3. Factor Loadings for Listening-Male

Item	Factor Loadings		
	1	2	3
1	0.585	0.005	-0.007
2	0.312	0.113	0.268
3	0.780	-0.013	0.004
4	0.402	0.072	0.035
5	0.442	0.148	0.023
6	0.625	0.084	-0.023
7	0.565	-0.047	-0.030
8	0.593	0.139	-0.121
9	0.641	0.149	0.145
10	0.267	-0.056	-0.201
11	0.308	0.862	-0.308
12	0.570	-0.162	-0.219
13	0.380	-0.128	0.084
14	0.440	0.018	0.197
15	0.581	0.097	0.199
16	0.585	-0.958	0.414
17	0.706	0.017	0.097
18	0.431	-0.175	-0.057
19	0.590	-0.056	-0.200

20	0.783	0.099	0.009
21	0.384	-0.286	-0.068
22	0.731	-0.229	-0.393
23	0.773	-0.082	-0.353
24	0.538	-0.378	-0.170
25	0.232	-0.130	-0.181
26	0.479	0.033	0.069
27	0.532	-0.011	0.057
28	0.692	-0.034	-0.112
29	0.645	0.117	-0.253
30	0.339	0.215	0.041
31	0.503	0.193	0.264
32	0.533	-0.279	-0.291
33	0.442	0.024	0.213
34	0.589	-0.316	0.505
35	0.288	0.115	0.106
36	0.569	0.049	0.073
37	0.418	0.238	0.040
38	0.287	0.019	0.110
39	0.049	-0.077	0.008
40	0.610	0.002	-0.178
41	0.340	0.209	0.225
42	0.345	0.149	0.184
43	0.405	-0.024	-0.198
44	0.462	0.099	-0.078
45	0.553	0.151	-0.170
46	0.391	0.218	0.343
47	0.452	0.071	0.036
48	0.247	0.130	0.130
49	0.312	0.416	0.367
50	0.289	0.207	0.151

Table A.4. Factor Loadings for Listening-Tagalog/Filipino

Item	Factor Loadings		
	1	2	3
1	0.598	-0.106	-0.138
2	0.399	-0.001	-0.147
3	0.686	-0.211	-0.053
4	0.452	0.148	0.147
5	0.448	-0.072	-0.267
6	0.556	-0.120	-0.146
7	0.663	-0.048	-0.220
8	0.532	0.061	-0.064
9	0.581	0.177	-0.261
10	0.321	-0.026	-0.043
11	0.216	0.023	-0.143

12	0.515	-0.121	0.122	4	0.473	-0.147	-0.198
13	0.376	0.181	-0.101	5	0.501	0.130	0.163
14	0.418	0.150	0.035	6	0.574	0.027	-0.159
15	0.541	0.114	-0.139	7	0.715	-0.025	-0.015
16	0.520	-0.392	-0.020	8	0.535	-0.013	-0.074
17	0.712	-0.034	-0.019	9	0.603	0.228	0.030
18	0.500	-0.115	0.083	10	0.374	0.176	0.017
19	0.526	-0.294	0.122	11	0.455	0.430	0.435
20	0.592	0.030	-0.083	12	0.712	-0.162	-0.080
21	0.421	0.127	-0.111	13	0.664	0.105	-0.206
22	0.668	-0.152	0.147	14	0.351	0.176	0.096
23	0.711	-0.160	-0.078	15	0.607	0.096	0.139
24	0.512	0.004	0.052	16	0.866	-0.167	-0.786
25	0.334	-0.291	0.266	17	0.722	-0.171	-0.110
26	0.581	-0.133	-0.120	18	0.747	-0.387	0.334
27	0.440	0.188	-0.273	19	0.675	-0.415	0.381
28	0.659	-0.301	0.078	20	0.719	-0.001	-0.042
29	0.693	-0.336	-0.049	21	0.459	-0.116	-0.212
30	0.345	-0.011	-0.268	22	0.480	-0.228	0.012
31	0.505	0.093	0.087	23	0.708	-0.062	0.031
32	0.539	-0.135	0.104	24	0.441	0.055	-0.216
33	0.459	0.153	-0.163	25	0.384	-0.176	0.501
34	0.514	0.385	-0.170	26	0.566	-0.110	-0.064
35	0.316	0.080	0.139	27	0.001	0.660	0.534
36	0.510	0.004	0.215	28	0.793	-0.493	0.420
37	0.424	0.266	0.092	29	0.666	-0.027	0.140
38	0.353	0.078	-0.036	30	0.322	0.128	0.191
39	0.097	0.146	0.237	31	0.214	0.025	-0.170
40	0.490	0.085	-0.047	32	0.535	0.154	-0.037
41	0.351	0.178	0.076	33	0.468	0.362	0.157
42	0.375	0.112	0.158	34	0.461	0.238	-0.063
43	0.444	0.120	0.389	35	0.431	0.010	0.192
44	0.530	0.114	0.203	36	0.554	-0.086	-0.097
45	0.456	0.004	0.337	37	0.412	-0.048	0.010
46	0.211	0.431	0.125	38	0.481	0.121	-0.470
47	0.455	0.312	0.096	39	0.145	-0.097	-0.128
48	0.263	0.006	0.286	40	0.353	0.298	-0.324
49	0.249	0.491	-0.025	41	0.333	0.429	0.339
50	0.376	0.207	0.182	42	0.288	0.345	0.034
				43	0.530	-0.183	-0.145
				44	0.507	0.091	0.126
				45	0.545	0.188	-0.015
				46	0.316	0.332	0.129
				47	0.548	0.048	-0.006
				48	-0.110	1.029	-0.467
				49	0.301	0.363	0.215
				50	0.322	0.140	0.024

Table A.5. Factor Loadings for Listening-Korean

Item	Factor Loadings		
	1	2	3
1	0.639	-0.044	-0.287
2	0.370	0.242	0.105
3	0.763	0.101	0.041

Table A.6. Factor Loadings for Listening-High Proficiency

Item	Factor Loadings		
	1	2	3
1	0.450	0.054	0.119
2	0.137	0.510	0.170
3	0.462	0.074	0.285
4	0.345	-0.144	-0.315
5	0.314	-0.057	0.407
6	0.477	-0.169	-0.004
7	0.145	0.018	0.139
8	0.128	0.175	0.067
9	0.022	0.169	-0.051
10	0.076	0.060	0.025
11	-0.026	0.586	-0.387
12	0.524	-0.489	-0.373
13	0.199	0.080	0.096
14	0.420	0.072	0.053
15	0.506	0.231	0.010
16	-1.498	0.316	-0.420
17	-0.024	0.056	-0.270
18	0.306	0.368	0.314
19	-0.133	-1.405	-0.043
20	-0.474	0.620	0.317
21	0.152	0.152	0.008
22	-0.346	-0.026	0.138
23	0.695	0.302	-0.493
24	-0.035	0.161	-0.095
25	0.282	-0.135	0.045
26	0.317	0.206	0.331
27	0.381	0.370	0.085
28	-0.179	0.059	1.388
29	0.463	0.380	-0.238
30	0.026	0.193	0.017
31	-0.034	0.459	0.066
32	0.335	-0.134	0.078
33	0.082	0.135	-0.018
34	0.133	0.188	0.188
35	0.128	-0.097	-0.009
36	-0.110	0.422	-0.451
37	-0.086	0.102	-0.008
38	0.022	0.061	-0.098
39	0.013	0.007	-0.028
40	-0.006	0.233	0.146
41	-0.154	-0.127	0.092
42	0.067	0.026	-0.028
43	-0.099	-0.075	0.238

44	-0.023	0.006	0.106
45	0.312	0.330	-0.352
46	0.087	-0.046	-0.010
47	-0.007	0.035	0.065
48	0.341	-0.027	-0.287
49	0.329	-0.109	0.146
50	-0.009	0.024	-0.014

Table A.7. Factor Loadings for Listening-Low Proficiency

Item	Factor Loadings		
	1	2	3
1	0.398	-0.273	-0.173
2	0.056	-0.145	0.212
3	0.575	-0.169	0.204
4	0.133	0.148	-0.068
5	0.143	-0.096	0.262
6	0.425	-0.348	0.079
7	0.444	-0.264	0.009
8	0.269	0.025	0.201
9	0.165	-0.210	0.276
10	0.139	-0.141	-0.153
11	0.142	0.067	0.378
12	0.529	-0.063	-0.179
13	0.085	-0.110	-0.186
14	0.183	0.242	0.189
15	0.266	0.013	0.260
16	0.636	-0.058	0.168
17	0.524	0.036	0.180
18	0.552	0.089	0.080
19	0.473	0.035	-0.052
20	0.449	0.060	0.036
21	0.216	-0.166	-0.314
22	0.579	-0.024	-0.080
23	0.564	-0.090	0.002
24	0.316	0.173	-0.220
25	0.293	0.122	0.017
26	0.405	0.028	0.023
27	0.168	-0.074	0.124
28	0.617	0.064	-0.173
29	0.601	-0.094	0.020
30	0.134	-0.156	0.004
31	0.231	0.079	0.228
32	0.485	-0.046	-0.175
33	0.136	0.000	-0.062
34	0.064	0.093	-0.168
35	0.182	0.113	0.085

36	0.446	0.324	0.071
37	0.011	0.238	-0.032
38	0.088	0.074	0.139
39	-0.037	0.134	-0.004
40	0.139	-0.173	-0.288
41	0.076	0.354	0.125
42	0.056	0.288	-0.071
43	0.361	0.295	-0.325
44	0.234	0.254	-0.227
45	0.322	0.439	0.045
46	-0.066	0.268	0.202
47	0.035	0.247	-0.158
48	0.069	0.105	-0.148
49	-0.190	0.114	0.191
50	0.044	0.139	0.014

Appendix B

Factor Loadings for Form EE-GCVR

Table B.1. Factor Loadings for GCVR -All Examinees

Item	Factor Loadings			
	1	2	3	4
1	0.553	-0.127	0.261	0.107
2	0.598	0.099	-0.012	-0.134
3	0.625	0.076	0.089	-0.002
4	0.629	-0.123	0.138	-0.101
5	0.382	-0.008	0.162	0.106
6	0.699	-0.028	0.186	-0.154
7	0.242	0.290	-0.003	-0.019
8	0.720	-0.033	0.257	0.032
9	0.323	0.281	-0.012	-0.085
10	0.457	0.346	-0.219	-0.177
11	0.649	-0.194	0.168	0.020
12	0.544	0.099	0.245	-0.140
13	0.593	-0.216	0.085	-0.032
14	0.268	0.503	0.060	-0.177
15	0.631	-0.069	-0.104	0.168
16	0.532	0.190	-0.057	-0.262
17	0.068	0.262	0.220	0.202
18	0.316	-0.243	0.343	0.488
19	0.026	0.628	-0.058	-0.091
20	0.203	0.171	-0.075	0.128
21	0.401	0.066	0.160	0.150
22	0.482	0.192	-0.007	-0.077
23	0.718	0.332	0.164	-0.080
24	0.546	0.281	0.022	-0.162
25	0.618	-0.107	0.129	-0.098
26	0.668	-0.161	-0.007	-0.265
27	0.585	0.265	0.371	-0.110
28	0.683	0.090	0.123	-0.017
29	0.559	0.165	0.188	0.015
30	0.492	0.434	-0.189	-0.183
31	0.448	0.406	0.115	-0.254
32	0.477	-0.059	-0.206	-0.055
33	0.228	0.040	0.222	0.088
34	0.545	0.172	0.136	0.001
35	-0.009	0.342	0.242	0.005
36	0.325	0.198	-0.091	0.002
37	-0.144	0.164	0.146	0.060
38	0.476	0.298	0.107	-0.124
39	0.251	0.371	-0.013	-0.249

40	0.585	-0.171	0.092	0.176
41	0.431	0.177	0.031	-0.163
42	0.554	0.017	0.038	-0.045
43	0.319	0.135	0.106	0.133
44	0.343	0.224	0.104	-0.015
45	0.274	0.189	-0.001	-0.052
46	0.365	-0.159	0.038	-0.120
47	0.430	0.309	-0.020	-0.039
48	0.258	0.369	0.130	0.085
49	0.675	-0.010	0.180	-0.164
50	0.277	0.209	0.146	0.111
51	0.716	-0.237	-0.235	-0.138
52	0.580	-0.389	-0.308	-0.088
53	0.567	-0.104	-0.257	-0.062
54	0.607	-0.163	-0.094	-0.102
55	0.478	0.193	-0.377	-0.259
56	0.501	-0.087	-0.085	-0.115
57	0.675	-0.522	0.112	0.030
58	0.641	-0.316	-0.056	0.137
59	0.743	-0.255	0.013	0.006
60	0.554	-0.183	-0.095	-0.186
61	0.637	-0.377	-0.098	-0.113
62	0.470	0.276	0.153	-0.080
63	0.382	0.070	0.046	-0.048
64	0.631	-0.309	-0.304	0.048
65	0.599	-0.145	-0.030	0.019
66	0.435	0.103	-0.153	-0.005
67	0.683	-0.196	-0.182	-0.001
68	0.651	-0.353	0.103	0.027
69	0.444	0.165	0.009	0.077
70	0.736	-0.263	-0.042	-0.106
71	0.610	0.034	-0.046	-0.085
72	0.617	-0.085	0.143	0.009
73	0.642	-0.287	0.172	-0.009
74	0.570	0.001	0.151	0.042
75	0.756	-0.329	-0.020	-0.026
76	0.719	-0.308	-0.178	0.137
77	0.566	-0.175	-0.018	0.077
78	0.633	-0.471	0.192	0.011
79	0.400	0.018	0.130	0.163
80	0.604	-0.126	0.027	-0.007
81	0.436	0.184	-0.038	0.190
82	0.421	0.160	-0.182	0.125

83	0.386	0.213	-0.255	0.319	25	0.602	-0.053	-0.186	-0.007
84	0.351	0.232	-0.228	0.279	26	0.704	-0.105	-0.070	-0.238
85	0.384	0.334	0.023	0.200	27	0.605	0.264	-0.407	-0.032
86	0.696	-0.169	-0.154	0.019	28	0.683	0.071	-0.129	0.020
87	0.393	0.189	-0.087	0.158	29	0.549	0.138	-0.192	0.095
88	0.511	0.329	-0.124	0.077	30	0.459	0.459	0.194	-0.179
89	0.593	-0.066	-0.155	0.168	31	0.388	0.470	-0.153	-0.314
90	0.420	0.384	-0.043	0.251	32	0.437	-0.027	0.132	-0.131
91	0.312	0.268	-0.128	0.261	33	0.188	0.026	-0.207	0.137
92	0.406	0.192	-0.067	0.273	34	0.536	0.135	-0.104	0.024
93	0.572	-0.038	-0.203	0.240	35	0.015	0.316	-0.209	-0.031
94	0.407	0.336	-0.079	0.275	36	0.342	0.209	0.136	0.012
95	0.227	0.255	-0.134	0.079	37	-0.191	0.117	-0.104	0.049
96	0.498	0.197	-0.163	0.116	38	0.436	0.366	-0.171	-0.098
97	0.084	0.275	-0.155	-0.015	39	0.242	0.462	-0.061	-0.267
98	0.435	0.085	-0.173	0.040	40	0.574	-0.213	-0.089	0.140
99	0.372	0.246	-0.198	0.179	41	0.447	0.158	-0.048	-0.095
100	0.565	0.093	0.002	0.157	42	0.531	0.026	-0.071	-0.080
					43	0.303	0.108	-0.052	0.215
					44	0.375	0.227	-0.122	-0.001
					45	0.274	0.244	0.004	-0.126
					46	0.379	-0.142	-0.063	-0.115
					47	0.416	0.334	-0.012	-0.011
					48	0.305	0.302	-0.070	0.112
					49	0.685	-0.002	-0.209	-0.076
					50	0.226	0.172	-0.161	0.156
					51	0.691	-0.242	0.276	-0.182
					52	0.599	-0.362	0.327	-0.135
					53	0.526	-0.042	0.287	-0.155
					54	0.566	-0.147	0.168	-0.186
					55	0.456	0.227	0.285	-0.192
					56	0.492	-0.085	0.101	-0.174
					57	0.647	-0.548	-0.084	-0.005
					58	0.603	-0.373	0.138	0.073
					59	0.708	-0.251	0.012	0.010
					60	0.539	-0.132	0.064	-0.253
					61	0.671	-0.344	0.078	-0.151
					62	0.453	0.241	-0.115	-0.107
					63	0.377	0.074	-0.002	-0.072
					64	0.596	-0.348	0.428	0.018
					65	0.599	-0.181	0.077	-0.005
					66	0.454	0.114	0.172	-0.055
					67	0.653	-0.170	0.218	-0.088
					68	0.606	-0.378	-0.133	-0.018
					69	0.455	0.165	0.005	0.051
					70	0.712	-0.208	0.046	-0.102
					71	0.607	0.040	0.079	-0.034
					72	0.594	-0.089	-0.173	0.083

Table B.2. Factor Loadings for GCVR-Female

Item	Factor Loadings			
	1	2	3	4
1	0.543	-0.158	-0.281	0.148
2	0.580	0.081	0.003	-0.152
3	0.608	0.068	-0.093	-0.029
4	0.615	-0.124	-0.248	-0.034
5	0.375	-0.005	-0.143	0.131
6	0.665	0.005	-0.252	-0.131
7	0.277	0.327	-0.012	-0.045
8	0.732	-0.106	-0.241	0.092
9	0.332	0.260	-0.001	-0.121
10	0.427	0.373	0.250	-0.219
11	0.612	-0.206	-0.220	0.046
12	0.554	0.101	-0.333	-0.053
13	0.598	-0.229	-0.140	-0.050
14	0.300	0.483	-0.085	-0.155
15	0.603	-0.092	0.091	0.145
16	0.564	0.157	0.035	-0.265
17	0.090	0.236	-0.149	0.255
18	0.343	-0.318	-0.103	0.414
19	0.046	0.615	0.079	-0.101
20	0.206	0.165	0.137	0.150
21	0.412	0.001	-0.065	0.184
22	0.478	0.178	0.095	-0.083
23	0.689	0.328	-0.195	-0.032
24	0.547	0.291	-0.054	-0.114

73	0.631	-0.356	-0.205	-0.023	16	0.450	0.268	0.150	-0.175
74	0.598	-0.026	-0.091	0.063	17	0.061	0.262	-0.008	-0.090
75	0.760	-0.346	0.025	-0.052	18	0.384	0.278	-1.207	0.049
76	0.668	-0.359	0.299	0.083	19	0.021	0.592	0.203	0.089
77	0.512	-0.224	0.083	0.080	20	0.210	0.126	0.192	-0.014
78	0.608	-0.536	-0.213	0.056	21	0.388	0.245	-0.108	-0.065
79	0.408	-0.007	-0.032	0.187	22	0.503	0.270	-0.080	-0.315
80	0.623	-0.124	-0.024	0.014	23	0.819	0.387	-0.224	-0.019
81	0.373	0.203	0.028	0.242	24	0.541	0.259	0.130	-0.138
82	0.387	0.182	0.192	0.103	25	0.644	-0.293	0.118	-0.250
83	0.380	0.227	0.330	0.265	26	0.584	-0.235	0.115	-0.215
84	0.334	0.222	0.287	0.296	27	0.559	0.242	-0.015	-0.300
85	0.431	0.304	0.002	0.166	28	0.687	0.178	0.003	-0.155
86	0.673	-0.132	0.136	0.061	29	0.571	0.212	0.365	-0.097
87	0.407	0.163	0.108	0.153	30	0.574	0.431	0.235	-0.149
88	0.456	0.374	0.073	0.052	31	0.645	0.291	-0.317	-0.316
89	0.558	0.028	0.195	0.243	32	0.570	-0.145	0.008	0.225
90	0.393	0.358	0.052	0.284	33	0.355	0.059	-0.095	-0.011
91	0.244	0.310	0.079	0.253	34	0.613	0.205	0.023	-0.068
92	0.354	0.216	0.081	0.266	35	-0.052	0.381	-0.181	0.041
93	0.510	0.041	0.168	0.262	36	0.304	0.094	0.067	-0.087
94	0.347	0.392	0.036	0.261	37	0.010	0.226	-0.072	-0.054
95	0.179	0.295	0.077	0.065	38	0.584	0.056	0.011	0.032
96	0.464	0.207	0.162	0.041	39	0.289	0.038	-0.013	-0.047
97	0.122	0.275	0.152	-0.054	40	0.685	-0.154	-0.071	0.227
98	0.382	0.108	0.142	0.037	41	0.400	0.172	0.179	-0.222
99	0.378	0.204	0.218	0.169	42	0.620	0.010	0.048	0.027
100	0.545	0.072	-0.024	0.199	43	0.379	0.180	0.248	-0.157
					44	0.267	0.094	0.137	0.182
					45	0.282	0.050	-0.114	0.089
					46	0.320	-0.187	-0.055	-0.054
					47	0.474	0.238	0.142	-0.044
					48	0.179	0.522	-0.051	0.077
					49	0.654	-0.082	0.066	-0.295
					50	0.423	0.266	0.147	-0.092
					51	0.745	-0.219	0.076	-0.206
					52	0.525	-0.410	0.009	-0.008
					53	0.643	-0.273	0.080	0.074
					54	0.702	-0.179	-0.033	0.059
					55	0.552	0.044	0.084	0.022
					56	0.530	-0.084	-0.103	0.116
					57	0.728	-0.457	-0.201	-0.055
					58	0.732	-0.183	-0.056	0.109
					59	0.802	-0.239	0.036	-0.074
					60	0.581	-0.326	0.006	0.096
					61	0.549	-0.473	-0.086	0.045
					62	0.526	0.409	-0.117	-0.149
					63	0.381	0.139	-0.142	-0.114

Table B.3. Factor Loadings for GCVR-Male

Item	Factor Loadings			
	1	2	3	4
1	0.587	-0.014	-0.171	0.059
2	0.639	0.170	0.177	-0.106
3	0.676	0.126	-0.038	0.011
4	0.648	-0.103	0.074	-0.046
5	0.408	-0.045	-0.123	-0.035
6	0.786	-0.112	-0.076	-0.079
7	0.157	0.138	-0.116	0.106
8	0.704	0.131	-0.034	-0.123
9	0.289	0.375	0.173	0.091
10	0.541	0.460	-0.363	-0.149
11	0.742	-0.089	0.028	-0.148
12	0.516	0.101	0.115	-0.003
13	0.582	-0.190	0.018	-0.042
14	0.187	0.543	0.192	-0.200
15	0.694	-0.020	0.169	0.157

64	0.687	-0.175	0.095	-0.120	6	0.611	-0.127	0.059	-0.043
65	0.605	-0.026	-0.071	0.179	7	0.294	0.032	-0.279	-0.228
66	0.396	-0.063	0.206	0.137	8	0.708	-0.062	0.116	-0.087
67	0.735	-0.267	-0.194	0.173	9	0.384	0.389	-0.033	-0.043
68	0.763	-0.300	-0.091	0.118	10	0.508	-0.040	-0.193	-0.458
69	0.438	0.078	-0.035	0.292	11	0.471	-0.139	0.149	-0.012
70	0.775	-0.365	0.015	-0.102	12	0.533	-0.099	0.019	-0.054
71	0.617	0.055	-0.053	-0.231	13	0.471	-0.008	0.146	-0.234
72	0.697	-0.197	0.127	-0.223	14	0.519	0.336	-0.280	0.032
73	0.696	-0.106	0.079	-0.161	15	0.521	-0.192	0.011	-0.032
74	0.491	0.149	-0.083	-0.079	16	0.516	-0.057	-0.029	-0.053
75	0.748	-0.325	-0.035	0.008	17	0.266	0.208	0.087	0.089
76	0.793	-0.158	0.028	0.006	18	0.357	0.066	-0.573	-1.042
77	0.676	0.034	0.062	-0.102	19	0.348	0.377	-0.197	0.056
78	0.679	-0.322	0.017	-0.066	20	0.242	-0.010	0.186	-0.014
79	0.411	0.106	-0.456	-0.205	21	0.303	-0.004	0.091	-0.078
80	0.572	-0.248	0.049	0.031	22	0.356	-0.147	-0.055	-0.176
81	0.602	0.118	0.067	-0.005	23	0.927	0.025	-0.084	-0.149
82	0.503	0.135	0.028	0.195	24	0.685	0.183	0.016	0.018
83	0.409	0.112	0.051	0.292	25	0.500	-0.368	-0.345	-0.006
84	0.430	0.174	0.261	0.203	26	0.600	-0.152	0.148	0.052
85	0.334	0.200	0.335	0.245	27	0.611	-0.151	-0.014	-0.031
86	0.751	-0.254	0.249	0.051	28	0.732	-0.206	0.111	-0.235
87	0.374	0.190	-0.017	0.301	29	0.565	-0.194	0.041	-0.013
88	0.686	0.222	-0.259	0.270	30	0.701	-0.055	0.106	-0.230
89	0.656	-0.230	-0.232	0.034	31	0.708	0.389	0.086	-0.079
90	0.504	0.322	0.165	0.154	32	0.253	-0.195	-0.337	0.371
91	0.471	0.181	-0.005	0.365	33	0.123	0.076	-0.396	-0.232
92	0.535	0.126	-0.081	0.329	34	0.616	-0.013	-0.045	-0.308
93	0.687	-0.239	0.026	0.297	35	0.014	0.605	-0.204	-0.049
94	0.575	0.116	-0.073	0.311	36	0.343	0.105	-0.070	0.228
95	0.349	0.191	0.069	0.160	37	-0.200	-0.001	-0.066	-0.003
96	0.604	0.181	-0.029	0.307	38	0.594	0.216	-0.092	0.122
97	0.002	0.236	0.128	0.241	39	0.370	0.364	-0.132	0.169
98	0.552	0.123	0.045	0.030	40	0.284	0.006	0.126	0.024
99	0.371	0.336	0.084	0.298	41	0.504	-0.074	0.145	-0.118
100	0.647	0.060	0.219	0.117	42	0.468	0.074	0.032	0.250
					43	0.357	0.019	0.043	0.051
					44	0.385	0.107	-0.030	0.248
					45	0.254	-0.082	0.121	-0.040
					46	0.285	-0.201	0.064	-0.032
					47	0.456	-0.164	0.075	-0.041
					48	0.406	0.270	0.292	-0.234
					49	0.600	-0.272	0.105	-0.079
					50	0.295	0.061	0.012	0.386
					51	0.524	-0.171	-0.175	0.500
					52	0.298	-0.309	-0.348	0.115
					53	0.351	-0.142	0.022	0.041

Table B.4. Factor Loadings for GCVR-Tagalog/Filipino

Item	Factor Loadings			
	1	2	3	4
1	0.434	-0.102	0.015	-0.089
2	0.534	-0.047	-0.015	-0.048
3	0.506	-0.203	0.076	-0.051
4	0.503	-0.163	0.161	-0.094
5	0.237	0.083	0.051	-0.037

54	0.459	0.088	-0.519	0.010
55	0.475	-0.234	0.123	-0.013
56	0.349	0.159	-0.057	-0.183
57	0.654	1.131	-0.367	0.677
58	0.574	-0.103	-0.046	0.291
59	0.550	-0.285	-0.219	0.178
60	0.388	0.013	0.080	-0.152
61	0.363	-0.223	-0.602	0.211
62	0.645	0.211	-0.082	0.028
63	0.310	0.171	0.057	0.100
64	0.133	0.658	1.330	-0.104
65	0.348	0.170	0.208	-0.121
66	0.384	-0.092	0.123	0.101
67	0.510	-0.088	0.291	0.380
68	0.120	-0.047	-0.259	0.436
69	0.494	-0.153	0.071	-0.010
70	0.637	-0.323	0.158	0.073
71	0.635	-0.246	0.107	0.018
72	0.626	-0.211	-0.001	0.156
73	0.462	0.039	0.336	0.448
74	0.531	-0.286	-0.022	-0.113
75	0.485	-0.079	0.202	0.050
76	1.022	0.866	-0.188	-0.471
77	0.410	0.000	-0.012	-0.117
78	0.435	0.092	0.328	-0.355
79	0.312	0.054	-0.101	0.017
80	0.486	-0.148	0.080	0.126
81	0.371	-0.075	-0.005	0.014
82	0.366	-0.097	0.140	0.162
83	0.389	-0.050	0.008	0.154
84	0.320	-0.099	0.047	0.057
85	0.438	-0.139	0.129	0.084
86	0.516	-0.120	0.098	0.108
87	0.449	0.098	0.028	0.108
88	0.502	-0.230	-0.093	0.076
89	0.427	-0.172	0.063	0.014
90	0.689	0.052	-0.326	-0.016
91	0.395	0.072	0.033	-0.005
92	0.365	-0.146	0.053	0.027
93	0.437	-0.202	-0.002	0.145
94	0.500	-0.116	0.077	-0.249
95	0.380	0.389	-0.021	0.059
96	0.575	0.223	0.016	0.210
97	0.204	0.454	-0.076	0.023
98	0.329	-0.212	0.175	0.143
99	0.455	-0.078	0.028	-0.161
100	0.580	-0.105	0.110	0.198

Table B.5. Factor Loadings for GCVR-Korean

Item	Factor Loadings			
	1	2	3	4
1	0.230	0.506	0.165	-0.187
2	0.644	0.022	0.131	0.145
3	0.466	0.184	-0.028	-0.240
4	0.477	0.289	0.031	0.034
5	0.137	-0.048	0.001	0.190
6	0.454	0.316	0.300	-0.020
7	0.309	0.289	-0.035	0.158
8	0.692	0.003	0.488	0.549
9	0.531	-0.053	-0.014	0.175
10	0.588	0.023	-0.101	-0.042
11	0.434	0.617	-0.108	-0.002
12	0.338	0.570	0.118	-0.118
13	0.280	0.276	-0.214	-0.211
14	0.161	0.076	-0.327	0.390
15	0.292	-0.196	0.132	0.164
16	0.354	-0.292	0.101	-0.292
17	0.020	0.084	-0.097	-0.474
18	0.121	-0.167	0.158	0.197
19	0.213	0.017	-0.095	-0.147
20	0.085	-0.517	-0.213	0.289
21	0.278	-0.328	0.037	-0.206
22	0.483	-0.187	-0.065	-0.178
23	0.249	-0.335	-0.069	0.039
24	0.524	-0.246	0.097	-0.135
25	0.316	0.415	0.097	0.055
26	0.532	0.299	0.298	0.338
27	0.307	0.706	0.463	0.149
28	0.330	-0.017	0.290	0.127
29	0.400	-0.154	0.120	-0.070
30	0.610	-0.004	0.215	-0.021
31	0.475	0.831	-0.284	-0.004
32	0.179	0.289	-0.384	0.212
33	-0.025	0.488	-0.075	-0.222
34	0.260	-0.294	0.004	-0.053
35	-0.058	-0.122	0.279	0.303
36	0.214	0.207	-0.228	0.078
37	-0.084	-0.079	-0.030	0.116
38	0.480	0.365	-0.036	-0.074
39	0.361	0.625	0.069	0.110
40	0.303	-0.095	0.019	0.268
41	0.330	0.476	0.046	-0.178
42	0.497	0.212	-0.014	0.104
43	0.252	-0.057	-0.191	0.317

44	0.083	-0.349	0.126	-0.129
45	0.274	0.249	0.172	-0.045
46	0.309	0.521	-0.232	-0.083
47	0.341	0.492	-0.314	0.157
48	0.278	0.146	0.052	0.128
49	0.436	0.773	-0.004	-0.027
50	-0.186	0.228	0.611	-0.212
51	0.588	-0.026	-0.229	-0.449
52	0.625	-0.505	0.101	-0.254
53	0.693	-0.431	-0.072	-0.134
54	0.439	0.246	-0.642	-0.113
55	0.540	-0.072	-0.318	-0.174
56	0.500	-0.304	0.322	-0.355
57	0.560	0.199	0.412	-0.285
58	0.357	-0.514	0.397	-0.155
59	0.454	-0.124	0.075	-0.455
60	0.543	0.078	-0.240	-0.537
61	0.559	0.301	-0.147	0.113
62	0.197	0.163	-0.082	-0.044
63	0.315	0.270	0.036	-0.043
64	0.428	-0.584	-0.188	0.038
65	0.529	-0.092	0.235	-0.098
66	0.451	-0.472	0.050	-0.199
67	0.343	-0.135	0.131	-0.081
68	0.425	-0.159	0.074	-0.095
69	0.204	-0.319	0.444	0.113
70	0.631	-0.095	0.061	0.106
71	0.451	-0.396	-0.096	-0.244
72	-0.175	-0.028	0.187	-0.019
73	0.271	-0.150	0.244	0.302
74	0.357	-0.116	-0.160	0.336
75	0.689	-0.203	0.018	0.174
76	0.365	-0.301	-0.283	0.056
77	0.352	-0.048	-0.295	-0.092
78	0.436	-0.292	0.097	-0.211
79	0.120	-0.284	0.274	-0.134
80	0.433	-0.159	0.389	-0.221
81	0.193	-0.289	-0.050	-0.040
82	0.295	0.139	0.313	0.160
83	0.214	-0.590	-0.325	0.413
84	0.474	-0.221	-0.355	0.031
85	0.233	-0.108	0.274	0.280
86	0.554	-0.202	-0.147	0.101
87	0.338	-0.148	-0.044	0.114
88	0.519	0.109	0.099	0.219
89	0.545	-0.261	-0.113	-0.044
90	0.241	-0.434	-0.272	0.333
91	0.426	-0.043	-0.434	0.217

92	0.531	-0.172	-0.234	0.012
93	0.567	0.033	-0.275	-0.029
94	0.475	0.017	0.233	0.250
95	0.242	0.272	0.148	0.267
96	0.546	-0.089	-0.022	0.170
97	0.214	0.101	-0.029	-0.042
98	0.294	0.061	-0.235	-0.008
99	0.334	-0.303	0.059	0.241
100	0.305	0.287	0.199	0.156

Table B.6. Factor Loadings for GCVR- High Proficiency

Item	Factor Loadings			
	1	2	3	4
1	0.355	0.024	-0.346	-0.145
2	0.303	0.184	-0.123	-0.064
3	0.404	0.128	-0.348	-0.059
4	0.390	-0.307	-0.421	-0.354
5	0.153	0.179	0.133	-0.331
6	0.565	-0.224	-0.269	0.106
7	0.179	0.271	-0.084	0.095
8	0.566	-0.145	0.106	-0.087
9	0.160	0.298	0.029	0.025
10	0.309	0.368	-0.146	0.101
11	0.323	-0.263	-0.028	-0.245
12	0.534	-0.013	-0.016	0.011
13	0.416	-0.227	0.040	-0.130
14	0.359	0.526	-0.130	-0.046
15	0.304	0.141	-0.164	-0.081
16	0.417	0.255	0.348	0.002
17	0.143	0.235	-0.114	-0.075
18	0.216	-0.265	-0.246	-1.134
19	0.118	0.725	0.066	0.049
20	0.101	0.134	0.230	-0.089
21	0.294	0.134	0.101	-0.180
22	0.277	0.270	-0.122	-0.055
23	0.693	0.275	0.303	-0.193
24	0.489	0.209	0.034	-0.013
25	0.385	-0.125	-0.007	-0.208
26	0.391	-0.482	-0.170	0.258
27	0.493	0.192	-0.079	-0.188
28	0.494	-0.022	0.115	-0.057
29	0.391	0.250	0.106	-0.211
30	0.393	0.371	-0.075	0.100
31	0.307	0.372	-0.205	0.169
32	0.182	0.115	-0.279	0.479
33	0.248	0.199	-0.036	-0.182

34	0.457	0.090	0.153	-0.092	82	0.157	0.344	-0.292	0.172
35	0.107	0.456	-0.033	-0.154	83	0.108	0.386	0.180	-0.157
36	0.176	0.213	-0.102	0.158	84	0.068	0.400	0.094	0.033
37	-0.025	0.257	-0.114	0.007	85	0.215	0.339	0.092	0.015
38	0.390	0.282	-0.115	0.069	86	0.355	-0.200	0.209	0.093
39	0.134	0.371	-0.090	0.205	87	0.213	0.204	0.276	0.078
40	0.336	-0.101	0.071	-0.074	88	0.283	0.417	0.080	0.071
41	0.380	0.181	-0.022	0.088	89	0.144	0.265	0.327	0.545
42	0.369	0.053	0.051	0.037	90	0.304	0.431	0.347	-0.172
43	0.172	0.148	0.093	-0.060	91	0.064	0.427	0.098	-0.042
44	0.310	0.167	0.053	-0.026	92	0.288	0.282	0.032	-0.048
45	0.167	0.246	-0.394	0.123	93	0.229	0.142	0.184	0.044
46	0.172	-0.201	-0.079	0.070	94	0.130	0.495	0.055	0.074
47	0.336	0.271	-0.308	0.220	95	0.109	0.285	-0.038	0.164
48	0.226	0.376	-0.018	-0.044	96	0.181	0.307	0.110	0.030
49	0.686	-0.219	-0.280	0.210	97	-0.036	0.299	-0.049	0.154
50	0.167	0.187	-0.061	-0.081	98	0.204	0.083	0.113	0.158
51	0.293	-0.159	-0.215	0.750	99	0.206	0.354	0.131	0.021
52	0.086	-0.534	0.341	0.614	100	0.411	0.109	0.122	0.002
53	0.320	-0.011	-0.176	0.533					
54	0.293	0.060	-0.525	-0.024					
55	0.137	0.311	0.007	0.118					
56	0.269	-0.073	0.154	0.012					
57	0.395	-0.562	-0.454	-0.374					
58	0.492	-0.469	0.297	-0.195					
59	0.532	-0.316	0.150	-0.190					
60	0.272	-0.171	0.023	0.180					
61	0.348	-0.424	0.135	0.160					
62	0.475	0.165	-0.082	-0.177					
63	0.324	0.221	-0.293	0.014					
64	0.095	-0.211	0.595	0.296					
65	0.377	-0.127	-0.438	0.048					
66	0.205	-0.006	0.083	0.064					
67	0.429	-0.239	0.411	0.445					
68	0.389	-0.244	0.107	-0.026					
69	0.507	0.108	0.407	0.031					
70	0.375	-0.432	0.215	0.193					
71	0.473	-0.109	0.250	0.037					
72	0.568	-0.304	0.188	-0.053					
73	0.565	-0.402	0.170	-0.036					
74	0.503	-0.110	0.011	-0.093					
75	0.477	-0.564	-0.176	0.122					
76	-0.122	0.144	1.653	-0.254					
77	0.453	-0.141	-0.319	0.110					
78	0.576	-0.615	0.291	-0.027					
79	0.326	0.096	0.019	-0.180					
80	0.453	-0.141	0.224	0.152					
81	0.187	0.298	0.052	-0.209					

Table B.7. Factor Loadings for GCVR-Low Proficiency

Item	Factor Loadings			
	1	2	3	4
1	0.407	-0.192	0.202	0.255
2	0.266	0.257	-0.105	0.120
3	0.234	0.053	0.055	0.036
4	0.405	-0.005	0.009	0.312
5	0.184	-0.158	0.119	-0.041
6	0.305	-0.023	-0.164	0.201
7	-0.081	0.194	0.195	0.016
8	0.324	-0.122	0.298	0.333
9	0.013	0.329	-0.060	0.198
10	0.078	0.523	-0.201	-0.040
11	0.518	-0.018	0.007	0.350
12	0.087	-0.080	0.030	0.399
13	0.456	-0.124	-0.042	-0.036
14	-0.157	0.291	-0.190	-0.059
15	0.464	0.147	0.108	0.041
16	0.129	0.262	-0.442	-0.062
17	-0.157	0.020	0.453	0.212
18	0.264	-0.217	0.342	0.005
19	-0.327	0.324	-0.017	0.099
20	-0.026	0.202	0.139	-0.142
21	0.157	-0.065	0.223	0.068
22	0.205	0.299	0.036	0.196
23	-0.104	0.349	-0.287	-0.090

24	0.018	0.356	-0.162	0.047	72	0.244	-0.086	-0.013	0.058
25	0.366	-0.076	-0.173	0.199	73	0.478	-0.280	-0.040	0.039
26	0.441	0.114	-0.238	0.271	74	0.185	-0.055	0.168	-0.003
27	-0.019	0.083	-0.073	0.433	75	0.730	-0.050	-0.043	-0.045
28	0.234	0.232	0.024	0.376	76	0.710	0.034	0.006	-0.154
29	0.120	0.032	-0.018	0.105	77	0.513	0.012	0.162	-0.157
30	-0.049	0.689	-0.133	0.014	78	0.622	-0.363	0.024	-0.057
31	-0.187	0.141	-0.236	0.405	79	0.220	-0.012	0.304	-0.145
32	0.377	0.134	-0.119	-0.023	80	0.432	-0.113	0.059	0.048
33	0.116	-0.195	0.091	0.132	81	0.087	0.151	0.120	-0.033
34	0.105	0.155	0.121	0.160	82	0.197	0.221	0.160	-0.043
35	-0.232	-0.147	0.089	0.024	83	0.126	0.350	0.249	-0.318
36	0.081	0.251	0.124	0.169	84	0.154	0.402	0.325	-0.083
37	-0.190	-0.166	0.076	-0.021	85	-0.032	0.272	0.331	0.125
38	-0.079	0.050	-0.178	0.031	86	0.549	0.058	0.009	-0.149
39	-0.112	0.264	-0.080	0.174	87	0.050	0.203	0.239	0.022
40	0.467	-0.136	0.170	-0.013	88	0.070	0.315	0.093	0.138
41	0.074	0.128	-0.129	-0.011	89	0.468	0.155	0.228	-0.080
42	0.284	0.005	-0.115	0.166	90	-0.061	0.427	0.198	-0.040
43	-0.003	-0.029	0.183	0.099	91	0.041	0.257	0.214	-0.029
44	0.022	0.171	0.190	-0.048	92	0.101	0.073	0.303	-0.301
45	0.039	0.118	0.118	0.009	93	0.437	0.202	0.189	-0.021
46	0.228	-0.171	-0.069	0.134	94	0.021	0.246	0.349	0.044
47	-0.007	0.358	0.033	0.221	95	-0.021	0.216	0.104	0.149
48	-0.154	0.177	0.242	-0.049	96	0.148	0.226	0.173	-0.119
49	0.279	0.042	0.003	0.381	97	-0.126	0.220	-0.049	-0.066
50	-0.022	0.161	0.134	0.329	98	0.126	0.195	-0.112	-0.045
51	0.687	0.158	-0.134	-0.119	99	0.076	0.260	0.100	-0.100
52	0.638	0.084	-0.164	-0.287	100	0.175	0.091	0.151	0.134
53	0.444	0.156	-0.101	-0.257					
54	0.524	0.070	-0.097	-0.003					
55	0.229	0.433	-0.267	-0.010					
56	0.317	0.002	-0.193	-0.159					
57	0.764	-0.358	-0.012	-0.065					
58	0.605	-0.004	0.113	-0.077					
59	0.659	-0.020	-0.028	-0.006					
60	0.495	0.036	-0.101	-0.103					
61	0.678	-0.080	-0.097	0.068					
62	-0.039	0.207	-0.059	0.032					
63	0.132	-0.020	-0.096	0.036					
64	0.673	0.166	-0.094	-0.127					
65	0.518	0.115	0.081	0.062					
66	0.176	0.360	0.092	-0.179					
67	0.578	0.107	-0.004	-0.092					
68	0.719	-0.205	0.012	0.079					
69	-0.030	0.025	0.139	-0.140					
70	0.644	0.006	-0.170	0.094					
71	0.213	0.138	-0.146	-0.209					